



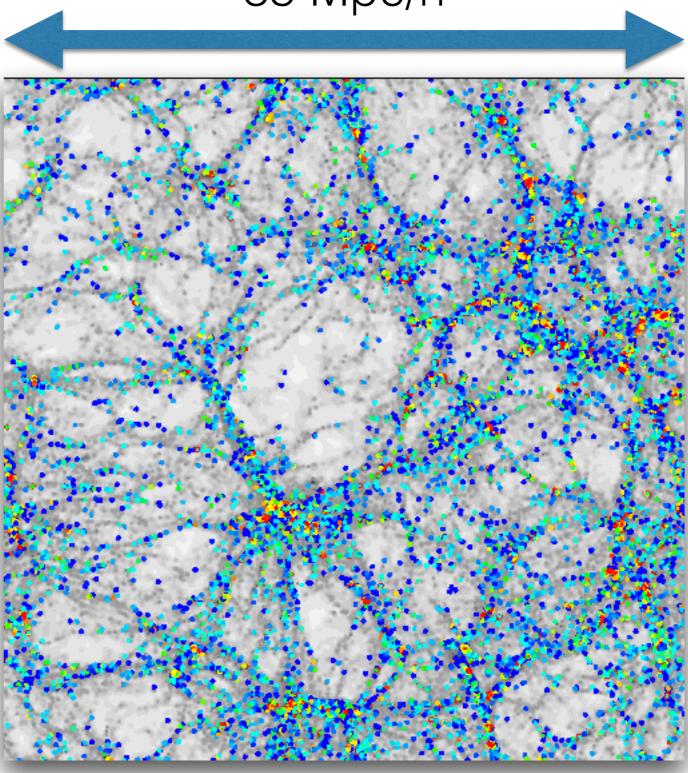
# Some insights into galaxy bias with a toy model

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## 85 Mpc/h

- simulation slice from z=0; LCDM
- gray: dark matter
- dots: B-V colors of galaxies



Credit: J. Colberg and A. Diaferio; GIF simulations (1998)

modes of density fluctuations (random fields):

$$\tilde{\delta}(\boldsymbol{k}) = \int \mathrm{d}^3 x \, \delta(\boldsymbol{x}) \, \mathrm{e}^{-\mathrm{i}\boldsymbol{x} \cdot \boldsymbol{k}}$$

complete second-order statistics of fluctuations:

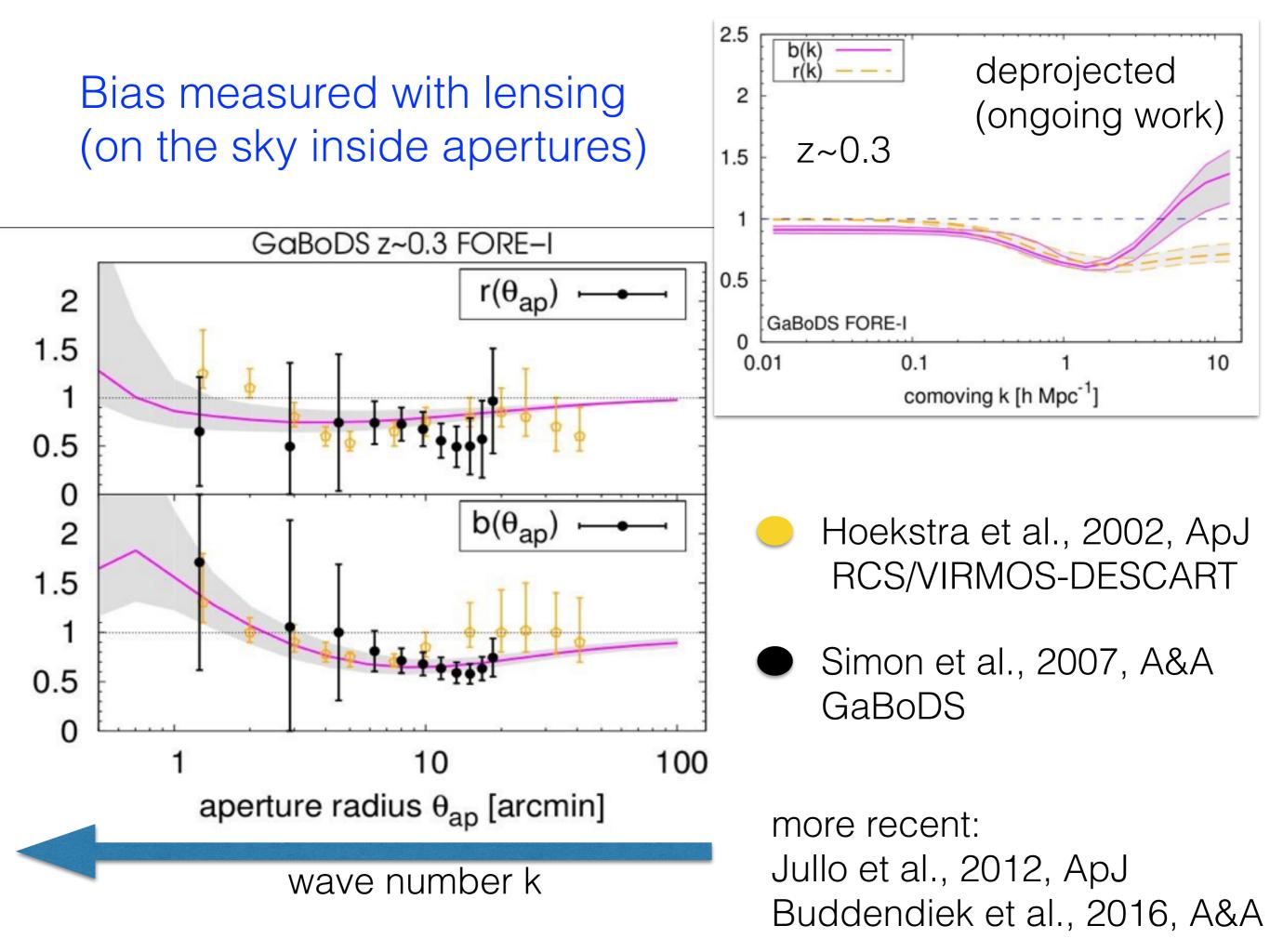
$$\begin{split} &\langle \tilde{\delta}_{\rm m}(\boldsymbol{k}) \tilde{\delta}_{\rm m}(\boldsymbol{k}') \rangle &= (2\pi)^3 \delta_{\rm D}(\boldsymbol{k} + \boldsymbol{k}') P_{\rm m}(\boldsymbol{k}) \ ; \\ &\langle \tilde{\delta}_{\rm m}(\boldsymbol{k}) \tilde{\delta}_{\rm g}(\boldsymbol{k}') \rangle &= (2\pi)^3 \delta_{\rm D}(\boldsymbol{k} + \boldsymbol{k}') P_{\rm gm}(\boldsymbol{k}) \ ; \\ &\langle \tilde{\delta}_{\rm g}(\boldsymbol{k}) \tilde{\delta}_{\rm g}(\boldsymbol{k}') \rangle &= (2\pi)^3 \delta_{\rm D}(\boldsymbol{k} + \boldsymbol{k}') \left( P_{\rm g}(\boldsymbol{k}) + \bar{n}_{\rm g}^{-1} \right) \ , \end{split}$$

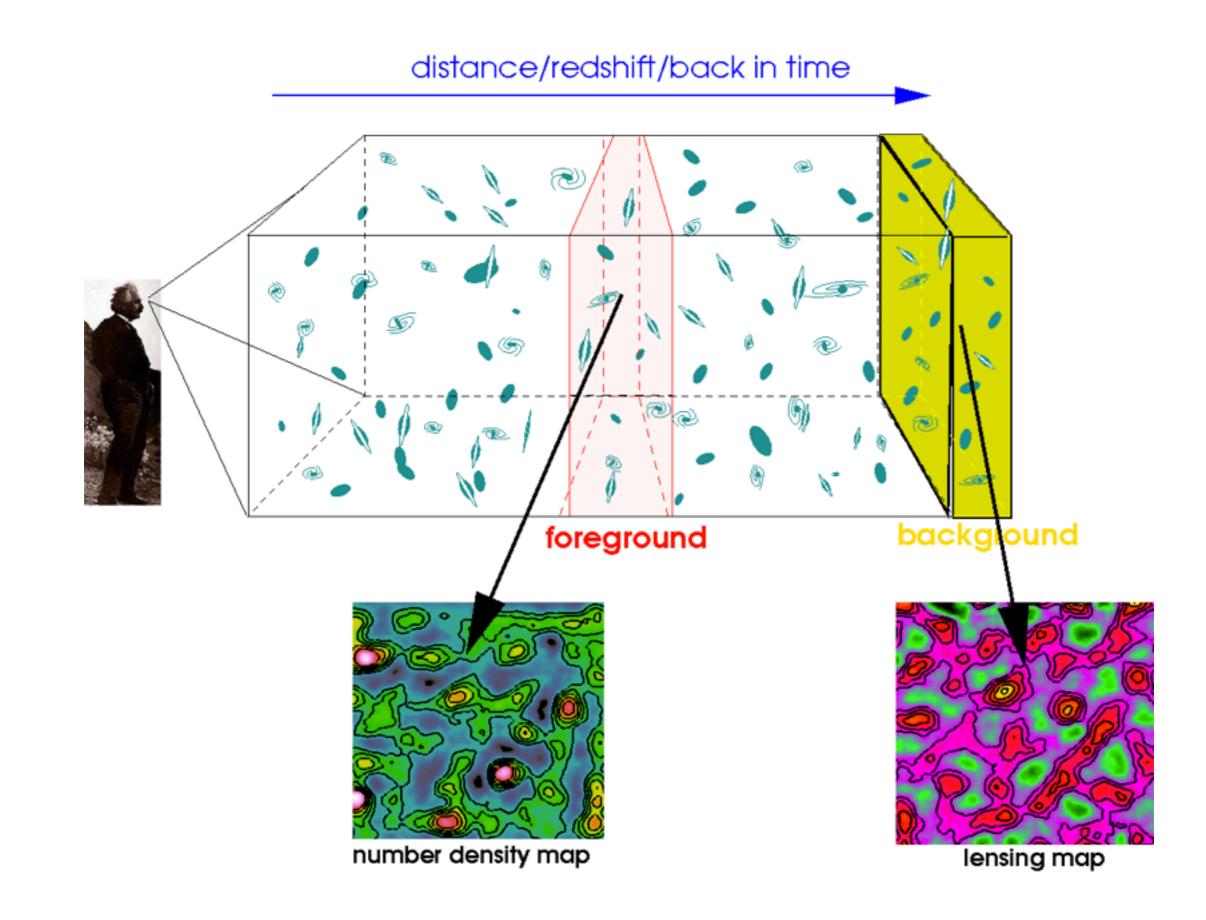
Biasing functions (linear stochastic bias):

bias factor

$$b(k) = \sqrt{\frac{P_{\rm g}(k)}{P_{\rm m}(k)}} \ ; \ r(k) = \frac{P_{\rm gm}(k)}{\sqrt{P_{\rm g}(k) \, P_{\rm m}(k)}} \ . \label{eq:bk}$$

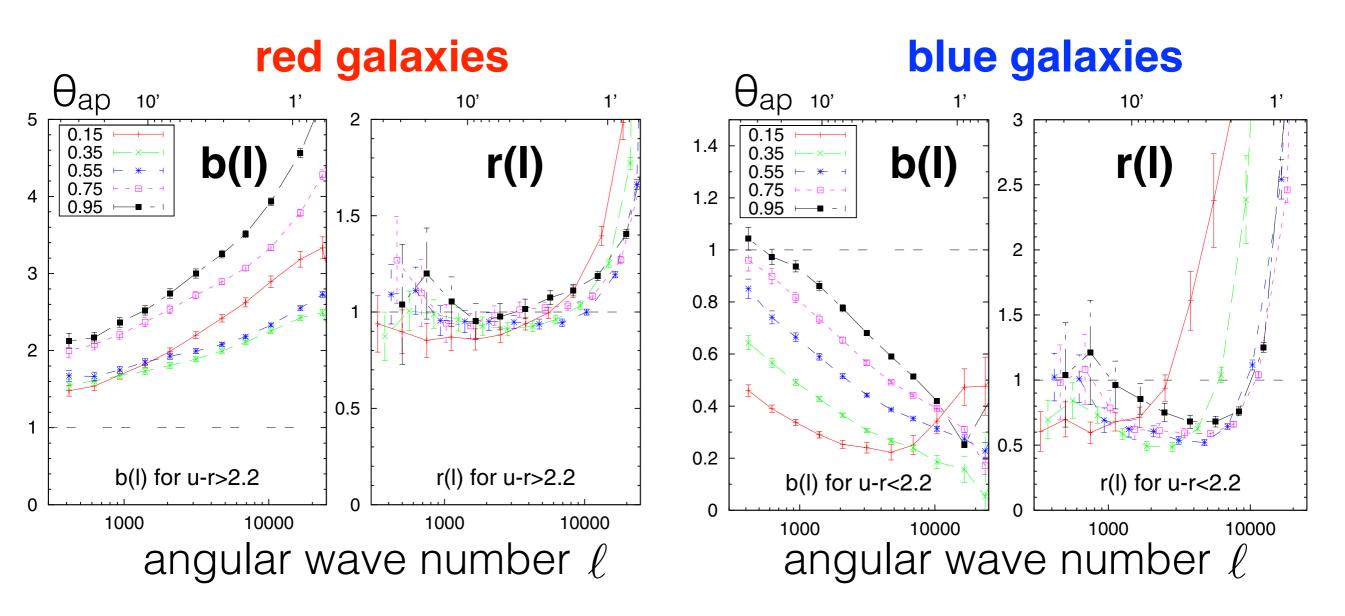
correlation factor





details: e.g. Simon et al., 2007, A&A, 861

In simulation: Millennium Simulation + SAMs (by colour and redshift; flux limit r < 25 mag)

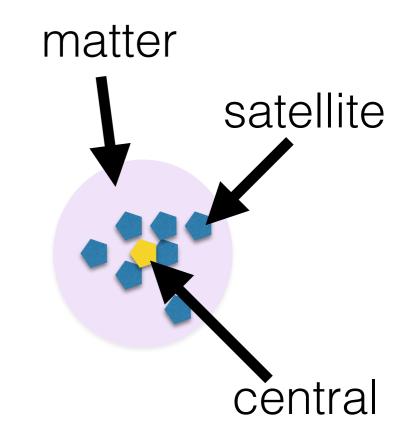


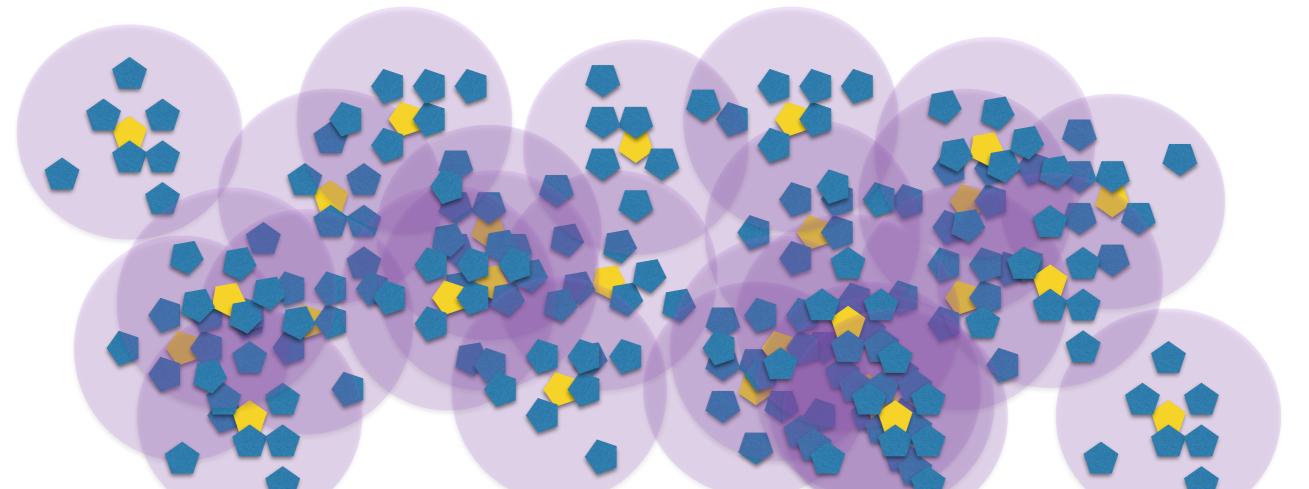
from Simon, P., 2013, A&A, 516

SAMs: Guo et al. (2011); dark matter: Springel et al. (2005)

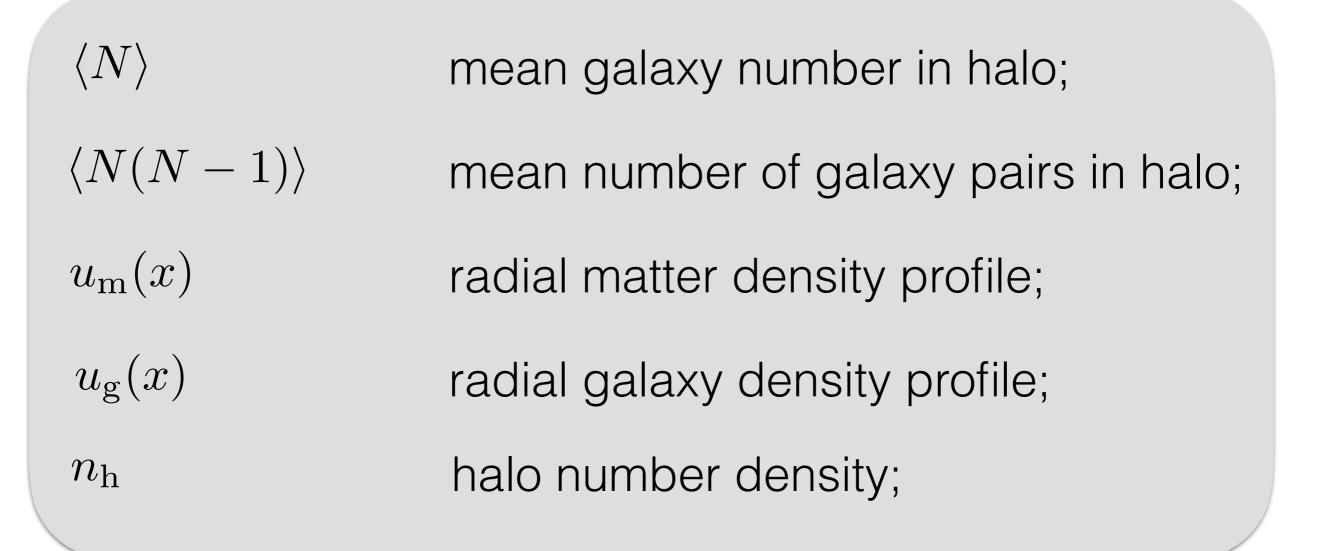
Halo-model inspired toy model that features:

- unclustered matter halos;
- single-mass halos with same density profile (mass m<sub>0</sub>);
- optionally mix of satellite and central galaxies or just satellite galaxies;





## Ingredients of model:



$$\tilde{u}(k) := \frac{\int_0^\infty \, \mathrm{d}x \, x \, k^{-1} \, u(x) \, \sin\left(kx\right)}{\int_0^\infty \, \mathrm{d}x \, x^2 \, u(x)}$$

Fourier transform of radial profile (normalised!) Implementation by Seljak, 2000, MNRAS, 218:

$$P_{g}(k) = \frac{n_{h}}{\bar{n}_{g}^{2}} \tilde{u}_{g}^{2p}(k) \left\langle N(N-1) \right\rangle = \frac{\left\langle N(N-1) \right\rangle}{n_{h} \left\langle N \right\rangle^{2}} \tilde{u}_{g}^{2p}(k)$$
$$P_{m}(k) = \frac{m_{0}^{2} n_{h}}{\bar{\rho}_{m}^{2}} \tilde{u}_{m}^{2}(k) = \frac{\tilde{u}_{m}^{2}(k)}{n_{h}}$$
$$P_{gm}(k) = \frac{n_{h} m_{0}}{\bar{\rho}_{m} \bar{n}_{g}} \tilde{u}_{m}(k) \tilde{u}_{g}^{q}(k) \left\langle N \right\rangle = \frac{\tilde{u}_{m}(k) \tilde{u}_{g}^{q}(k)}{n_{h}}$$

uses  $\bar{\rho}_{\rm m} = n_{\rm h} m_0$ ,  $\bar{n}_{\rm g} = n_{\rm h} \langle N \rangle$ ,  $n(m) = n_{\rm h} \delta_{\rm D}(m - m_0)$ , and  $p = \begin{cases} 1 & \langle N(N-1) \rangle > 1 \\ 1/2 & \text{otherwise} \end{cases}$   $q = \begin{cases} 1 & \langle N \rangle > 1 \\ 0 & \text{otherwise} \end{cases}$ 

controls impact of central galaxies

Correlation factor for scale *k*:

$$r(k) = \frac{P_{\rm gm}(k)}{\sqrt{P_{\rm g}(k) P_{\rm m}(k)}} = \frac{\tilde{u}_{\rm g}^{q-p}(k) \langle N \rangle}{\sqrt{\langle N(N-1) \rangle}} = \tilde{u}_{\rm g}^{q-p}(k) \left(1 + \frac{\Delta \sigma_{\rm N}^2}{\langle N \rangle^2}\right)^{-1/2}$$

where we have introduced the excess variance

$$\Delta \sigma_{\rm N}^2 := \sigma_{\rm N}^2 - \sigma_{\rm N}^2 |\text{Poisson}$$
$$= \langle N^2 \rangle - \langle N \rangle^2 - \langle N \rangle = \langle N(N-1) \rangle - \langle N \rangle^2$$

 $\Delta \sigma_{\rm N}^2 = 0$  variance of N as in Poisson statistic;

- $\Delta \sigma_{\rm N}^2 < 0$  sub-Poisson variance;
- $\Delta \sigma_{\mathrm{N}}^2 > 0$  super-Poisson variance

$$r(k) = \tilde{u}_{g}^{q-p}(k) \left(1 + \frac{\Delta \sigma_{N}^{2}}{\langle N \rangle^{2}}\right)^{-1/2}$$

We learn about two ways to produce r(k) > 1:

1. dominating central galaxies; for q-p = -1/2, r(k) can become arbitrarily large for k >> 1:

$$r(k) \propto \tilde{u}_{\rm g}^{-1/2}(k)$$

2. no/negligible central galaxies and sub-Poisson variance; hence p = q = 1 and  $\Delta \sigma_N^2 < 0$ ;

Note: impact of 2. becomes small if  $\Delta \sigma_N^2 << <N>^2$ 

### Bias factor for scale *k*:

$$b(k) = \sqrt{\frac{P_{\rm g}(k)}{P_{\rm m}(k)}} = \frac{\tilde{u}_{\rm g}^p(k)\sqrt{\langle N(N-1)\rangle}}{\tilde{u}_{\rm m}(k)\langle N\rangle} = \frac{\tilde{u}_{\rm g}^q(k)}{\tilde{u}_{\rm m}(k)}\frac{1}{r(k)}$$

- even without central galaxies and identical radial profiles we do not necessarily have b(k) = r(k) = 1; yet we find b(k)r(k) = 1;
- only a Poisson variance in addition to 1. ensures
   b(k) = r(k) = 1; truly unbiased galaxies require a Poissonlike variance of N;
- 3. for satellite-dominated halos, b(k) reflects the difference between matter and galaxy radial profile; hence q = 1 and r(k) = const;

- Central galaxies or a sub-Poisson variance of N can make r(k) > 1 because
  - P<sub>g</sub>(k) is defined in excess of Poisson shot noise of discrete galaxies. If galaxy sampling is actually sub-Poisson, this over-corrects the shot noise power (Seljak 2000; Guzik & Seljak 2001).
  - 2. Putting one galaxy *always* at the halo centre is not a Poisson sampling of the halo density profile; so is non-Poisson  $\Delta \sigma_N^2$ .
- Is r(x) > 1 also possible for the real-space biasing functions (shot noise only at zero lag)?

$$b(x) = \sqrt{\frac{\xi_{g}(x)}{\xi_{m}(x)}}; r(x) = \frac{\xi_{mg}(x)}{\sqrt{\xi_{g}(x)\xi_{m}(x)}}$$

-> transform three power spectra of toy model to real-space correlation functions via (back Fourier transform)

$$\xi(x) = [P](x) := \frac{1}{2\pi^2 x} \int_0^\infty dk \, k \, P(k) \, \sin(k \, x) \,,$$

and you get real-space counterparts of b(k) and r(k):

$$b(x) = \frac{[\tilde{u}_{\rm m} \cdot \tilde{u}_{\rm g}^q](x)}{|[\tilde{u}_{\rm m}^2](x)|} \frac{1}{r(x)};$$
  

$$r(x) = \frac{[\tilde{u}_{\rm m} \cdot \tilde{u}_{\rm g}^q](x)}{\sqrt{[\tilde{u}_{\rm g}^{2p}](x)[\tilde{u}_{\rm m}^2](x)}} \left(1 + \frac{\Delta \sigma_{\rm N}^2}{\langle N \rangle^2}\right)^{-1/2}$$

Interestingly, r(x) depends also on matter density profile.

Assume: identical density profiles  $u_m(x) = u_g(x)$ 

• without central galaxies — p = q = 1 — we have

$$b(x) r(x) = 1 ; r(x) = \left(1 + \frac{\Delta \sigma_{\mathrm{N}}^2}{\langle N \rangle^2}\right)^{-1/2}$$

• with centrals dominating and Poisson variance -p = 1/2, q = 0, and  $\Delta \sigma_N^2 = 0$  — we have

$$b(x) r(x) = \frac{[\tilde{u}_{\rm m}](x)}{|[\tilde{u}_{\rm m}^2](x)|} ; \ r(x) = \sqrt{\frac{[\tilde{u}_{\rm m}](x)}{[\tilde{u}_{\rm m}^2](x)}} \ge 1$$

# Summary and conclusions

- biasing functions of linear stochastic bias fully capture the differences in distributions of galaxies vs. matter for two-point statistics; can be measured with lensing;
- their definition assumes Poisson shot-noise for galaxies; can give seemingly curious values r > 1 if sampling is actually sub-Poisson;
- demonstrated with toy model that r(k) > 1 or r(x) > 1 can arise through central galaxies or a sub-Poisson variance of galaxy numbers inside matter halos;
- even if galaxies perfectly trace matter, we still have not b = r = 1 for all scales if sampling variance is not Poisson;